

Time evolution towards q -Gaussian stationary states through unified Itô-Stratonovich stochastic equation

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(Received 13 April 2010; published 13 December 2010)

We consider a class of single-particle one-dimensional stochastic equations which include external field, additive, and multiplicative noises. We use a parameter $\theta \in [0, 1]$ which enables the unification of the traditional Itô and Stratonovich approaches, now recovered, respectively, as the $\theta=0$ and $\theta=1/2$ particular cases to derive the associated Fokker-Planck equation (FPE). These FPE is a *linear* one, and its stationary state is given by a q -Gaussian distribution with $q = \frac{\tau+2M(2-\theta)}{\tau+2M(1-\theta)} < 3$, where $\tau \geq 0$ characterizes the strength of the confining external field and $M \geq 0$ is the (normalized) amplitude of the multiplicative noise. We also calculate the standard kurtosis κ_1 and the q -generalized kurtosis κ_q (i.e., the standard kurtosis but using the escort distribution instead of the direct one). Through these two quantities we numerically follow the time evolution of the distributions. Finally, we exhibit how these quantities can be used as convenient calibrations for determining the index q from numerical data obtained through experiments, observations, or numerical computations.

DOI: [10.1103/PhysRevE.82.061119](https://doi.org/10.1103/PhysRevE.82.061119)

PACS number(s): 02.50.-r, 05.20.-y, 05.40.-a, 05.90.+m

I. INTRODUCTION

The random walk is the simplest model of diffusive processes in physics. If there is no wind, or any other source of symmetry breaking, a drunk has a probability of 1/2 to take a step to the right and the same probability to take a step to the left at each instant. Underneath this model there is a formal mathematical theory known as stochastic calculus, which in some sense can be interpreted as an extension of the standard differential calculus taught in undergraduate courses. The stochastic as well as the standard calculus is based on the definition of the integral. Let us define the stochastic integral by

$$I[G(t)] = \int_{t_0}^t dW(t')G(t'), \quad (1)$$

where $G(t)$ is a left-continuous function (i.e., a function which is continuous from the left at all the points where it is defined) and $W(t)$ is a Wiener process [1,2]. As in the Riemann integral definition, the formal stochastic integral (also known as Riemann-Stieltjes integral) is a infinite discrete sum of very small intervals $[dW(t')]$ of a stochastic function. When we perform this sum we must make a choice; more precisely the function $G(t')$ has to be evaluated in some point inside each interval. This choice defines what kind of stochastic calculus will be performed henceforth. The two most famous procedures are Itô calculus and Stratonovich calculus. In the first, Itô calculus, the function is evaluated at the beginning of the intervals:

$$I_I[G(t)] = \text{ms-lim} \sum_{n \rightarrow \infty} G_{i-1} \Delta W_i, \quad (2)$$

where ms-lim stands for *mean-squared* limit [1], i.e., a second moment convergence; n is the number of subintervals in which we divide the interval $[t_0, t]$; and $\Delta W_i = W_i - W_{i-1}$. In the Stratonovich calculus we take the arithmetic average between the integrate function values at the beginning and at the end of the intervals as follows:

$$I_S[G(t)] = \text{ms-lim} \sum_{n \rightarrow \infty} \frac{G_i + G_{i-1}}{2} \Delta W_i. \quad (3)$$

A unified form, with a linear interpolation, can be proposed [3–8] for the stochastic integral, namely,

$$I_\theta[G(t)] = \text{ms-lim} \sum_{n \rightarrow \infty} [\theta G_i + (1 - \theta)G_{i-1}] \Delta W_i, \quad (4)$$

where $0 \leq \theta \leq 1$. Notice that the two traditional procedures can be recovered easily. Indeed, if $\theta=0$ we obtain the Itô approach, and if $\theta=1/2$ we obtain the Stratonovich approach. Moreover, if $\theta=1$ we recover the so-called backward-Itô [9] stochastic approach, also known as isothermal convention [3–5,7,8], or even as kinetic form [10,11].

We argue here that, in fact, it is possible to go one step further and generalize Eq. (4). Indeed, we can assume that the values of θ , at each interval $dW(t')$, are given by an arbitrary distribution $\rho(\theta)$ ($\int_0^1 d\theta \rho(\theta) = 1$). In particular, if this distribution is $\rho(\theta) = \delta(\theta - \theta_0)$, we recover the three cases mentioned above for suitable values of θ_0 , namely, $\theta_0=0$ (Itô), $\theta_0=1/2$ (Stratonovich), and $\theta_0=1$ (backward Itô). An interesting remark arises when $\rho(\theta)$ is constant. In this case $\langle \theta \rangle = 1/2$, which coincides precisely with the average value corresponding to the (frequently adopted) Stratonovich approach. In some sense, this is what seems to happen in most experiments, where the act of measuring is not instantaneous but in a time window. It is argued in [12] that the Itô particular case is the only one to be strictly consistent with causality. However, from a different viewpoint, it is argued

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in [13] that Itô calculus cannot provide a probability density function for the energy of a Brownian particle, in contrast with the Stratonovich calculus which correctly satisfies this requirement. Moreover, it is argued in [10,11] that only the backward-Itô approach, i.e., the kinetic form, is the only one to be consistent with the generalized Einstein relation. As we see from this diversity of statements, the problem still has open points. However, this discussion does not constitute the main scope of the present paper. Consistently, we address the generic case, i.e., for arbitrary values of θ (which could in fact correspond to different experimental realizations). Therefore, based on definition (4), we calculate the following integral:

$$\begin{aligned} I_{\theta}[W(t)] &\equiv \int_{t_0}^t dW(t')W(t') \\ &= \text{ms-lim}_{n \rightarrow \infty} \sum_{i=1}^n [\theta W_i + (1-\theta)W_{i-1}] \Delta W_i \\ &= \frac{1}{2} \{ [W(t)]^2 - [W(t_0)]^2 - (1-2\theta)(t-t_0) \}. \end{aligned} \quad (5)$$

Let us remind here that, in this equation, the limits $\theta=0$, $\theta=1/2$, and $\theta=1$, respectively, correspond to the Itô, Stratonovich, and kinetic approaches. For more details about the derivation of the last equation, see the Appendix.

In addition to the ingredients of stochastic calculus that we have mentioned above, some other concepts will be necessary in the present paper. In particular, escort mean values [14,15] are convenient theoretical tools for describing basic features of some probability densities, mainly those densities which decay as power laws that naturally appear in the study of complex systems dynamics, such as those obeying nonextensive statistical mechanics [16]. This theory generalizes the Boltzmann-Gibbs (BG) statistical mechanics and is governed by an entropic index q , which equals unity for the BG case. The characterization of a probability density by its set of escort mean values, if all of them converge, is a natural extension of the well-known characterization of a distribution in terms of its standard moments, which corresponds to $q=1$. The q -generalized theory has been applied to calculate many features of several complex systems, such as (i) the velocity distribution of cells of *Hydra viridissima* that follows a $q=3/2$ probability density function (PDF) [17]; (ii) the velocity distribution of (cells of) *Dictyostelium discoideum* that follows a $q=5/3$ PDF in the vegetative state and a $q=2$ PDF in the starved state [18]; (iii) the velocity distribution in defect turbulence [19]; (iv) the velocity distribution of cold atoms in a dissipative optical lattice [20]; (v) velocity distribution during silo drainage [21,22]; (vi) the velocity distribution in a driven-dissipative two-dimensional dusty plasma, with $q=1.08 \pm 0.01$ and $q=1.05 \pm 0.01$ at temperatures of 30 000 and 61 000 K, respectively [23]; (vii) the spatial (Monte Carlo) distributions of a trapped $^{136}\text{Ba}^+$ ion cooled by various classical buffer gases at 300 K [24]; (viii) the distributions of price returns at the stock exchange [25–27]; (ix) the distributions of returns of magnetic field fluctuations in the solar wind plasma as observed in data

from Voyager 1 [28] and from Voyager 2 [29]; (x) the distributions of returns of the avalanche sizes in the Ehrenfest dog-flea model [30]; (xi) the distributions of returns of the avalanche sizes in the self-organized Olami-Feder-Christensen model, as well as in real earthquakes [31]; (xii) the distributions of angles in the hamiltonian mean field (HMF) model [32]; (xiii) the distribution of stellar rotational velocities in the Pleiades [33]; (xiv) the distribution of transverse momenta in high-energy proton-proton collisions [34]; and (xv) in the action of spin glasses [35]. In fact, before the q -generalized theory had been proposed, some of its associated distributions were already found (with different names, naturally) in various fields of physics, for example, (i) the so-called κ distribution (for velocities) in plasma physics [36]; (ii) the well known t distribution [37], which is a q -Gaussian distribution with particular values of q ; and (iii) the famous Cauchy-Lorentz distribution, which is a $q=2$ distribution and appears frequently in many branches of physics, particularly in optics [38].

In the present work we use Eq. (5) to write unified Langevin and Fokker-Planck equations (FPEs). In Sec. II, after deriving these equations, we will show that the steady solutions of the FPE are q Gaussians (see later on for their precise definition) whose entropic parameter q depends on the noise and drift amplitudes, and also on the choice of a stochastic approach represented by the parameter θ . In Sec. III we introduce a generalized kurtosis based on [15] to characterize the temporal evolution of the distribution. After that, we integrate numerically the FPE with an initial distribution different from its asymptotic form. In particular, we consider as initial distributions q Gaussians characterized by an index q_i ($q_i \neq q$). Our numerical results show how the convergence towards the attractor behaves as a function of the parameter θ . Finally, we compare the standard kurtosis ($q=1$) with the one calculated using escort mean values, namely, q kurtosis. Our results show that the standard kurtosis has a divergence at $q=7/5$, while the q kurtosis has no divergence in the range $-1 < q \leq 3$. In addition to that, we show that the standard kurtosis and the q kurtosis are monotonic functions of the entropic parameter q , which suggests that they could be used as calibration curves to determine, from numerical data, the most appropriate value of q . Let us now describe the consequences of the unification on the Langevin and Fokker-Planck descriptions.

II. FOKKER-PLANCK EQUATION AND ITS SOLUTIONS

A stochastic differential equation (SDE) is not completely defined by itself. If the Langevin equation has multiplicative noise, we must choose what approach, Itô or Stratonovich, will be used to integrate it. This is the well-known Itô-Stratonovich dilemma [1–3] and is ultimately solved by taking into account the specific features of the system under investigation. For instance, if the noise has a finite correlation time τ_c (even if the limit $\tau_c \rightarrow 0$ is used to derive the SDE) or the noise comes from external sources, the Stratonovich choice is the adequate one. Instead, the Itô formalism is the correct choice if τ_c is strictly zero or the noise

comes from internal sources [1,39]. For external noise sources we mean those which have a parameter that permits one, in principle, to turn off the noise or even those that are not influenced by the system itself. In contrast, internal noise sources are those which fluctuations are due to inherent parts of the mechanism by which the state of the system evolves. This kind of noise cannot be turned off by manipulating a parameter [40]. A very interesting and enlightening discussion about the Stratonovich-Itô dilemma can be found in [40]. This controversy between the two approaches appears in the form of the so-called noise-induced drift which is an effect of the state dependence of the noise strength.

Based on proposal (4) we will write an unified Fokker-Planck equation that contains, as special cases, the Itô ($\theta=0$) and Stratonovich ($\theta=1/2$) forms. To do it, let us consider the quite general SDE for a stochastic variable $u(t)$:

$$du(t) = a[u(t), t]dt + b[u(t), t]dW(t), \quad (6)$$

where $a[u(t), t]$ is a deterministic external force given, for instance, by a potential, $b[u(t), t]$ is the state-dependent noise amplitude, and $dW(t) = \xi(t)dt$ is a Wiener increment [$\xi(t)$ is a Gaussian zero-mean white noise]. We will use a shorter notation for the dynamic variable $u(t)$ omitting its time dependence, so from now on $u \equiv u(t)$, $a[u(t), t] \equiv a(u, t)$, and $b[u(t), t] \equiv b(u, t)$, unless it is indispensable for the clarity of the text.

We are interested in the dynamics of the distribution function of u , so once we have the Langevin equation we may obtain a FPE for the probability density $P(u, t)$ by the Kramers-Moyal expansion [2] $\partial_t P = \sum_{n \geq 1} (-\partial_u)^n [D^{(n)} P]$, where the coefficients $D^{(n)}$ are given by

$$D^{(n)}(x, t) = \frac{1}{n!} \lim_{\epsilon \rightarrow 0} \frac{\langle [u(t+\epsilon) - x]^n \rangle}{\epsilon} \Big|_{u(t)=x}. \quad (7)$$

To calculate these coefficients, we need to write the Langevin equation in the integral form

$$u(t+\epsilon) - x = \int_t^{t+\epsilon} dt' a[u(t'), t'] + \int_t^{t+\epsilon} dW(t') b[u(t'), t'], \quad (8)$$

and assume that $a(u, t)$ and $b(u, t)$ can be expanded as

$$a[u(t+\epsilon), t+\epsilon] = a[x, t+\epsilon] + a'[x, t+\epsilon]\epsilon + \dots, \quad (9)$$

$$b[u(t+\epsilon), t+\epsilon] = b[x, t+\epsilon] + b'[x, t+\epsilon]\epsilon + \dots, \quad (10)$$

where a' and b' means differentiation with respect to u . Putting Eqs. (9) and (10), up to first order on ϵ , into Eq. (8) and iterating the result, we get

$$\langle u(t+\epsilon) - x \rangle = \epsilon a(u, t+\epsilon) + b'(u, t+\epsilon) b(u, t+\epsilon) \times \left\langle \int_t^{t+\epsilon} W(t') dW(t') \right\rangle \quad (11)$$

$$= \epsilon [a(u, t+\epsilon) + \theta b'(u, t+\epsilon) b(u, t+\epsilon)]. \quad (12)$$

If the noise in the Langevin equation is δ correlated, after repeated iterations it is possible to show that, at ϵ order, the terms of the Kramers-Moyal expansion with $n \geq 3$ vanish. Using this arguments we can calculate the first two coefficients of that expansion, namely,

$$K_\theta(u, t) \equiv D^{(1)}(u, t) = a(u, t) + \theta b'(u, t) b(u, t), \quad (13)$$

$$D(u, t) \equiv D^{(2)}(u, t) = b^2(u, t). \quad (14)$$

This procedure leads to the following Fokker-Planck equation:

$$\frac{\partial P}{\partial t} = - \frac{\partial}{\partial u} \left\{ K_\theta(u, t) P - \frac{1}{2} \frac{\partial}{\partial u} [D(u, t) P] \right\}. \quad (15)$$

We recover, for $\theta=0$ and $\theta=1/2$, respectively, the Itô and Stratonovich forms currently found in the literature. As expected, the difference between the two forms appears in the drift term K_θ , not in the diffusion term. If we deal with a FPE in the Stratonovich form, we should include the above-mentioned noise-induced drift term $\frac{1}{2} \partial_u [D(u)]$, which is unnecessary in the Itô form. We can write these FPE as a continuitylike equation $\partial_t P = -\partial_u j(u)$ if we define $j(u) = K_\theta(u) P - \frac{1}{2} \partial_u [D(u) P]$ as a probability current. In this case, the stationary solution of the FPE can be obtained elegantly from zero-flux boundary conditions $j(\pm\infty) = j(u) = 0$. It is very interesting to notice that this probability current can be rewritten as $j(u) = K_0(u) P - \frac{1}{2} [D(u)]^{2\theta} \partial_u \{ [D(u)]^{1-2\theta} P \}$, which allows us to rewrite the FPE as

$$\frac{\partial P}{\partial t} = - \frac{\partial}{\partial u} \left\{ K_0(u) P - \frac{1}{2} [D(u)]^{2\theta} \frac{\partial}{\partial u} \{ [D(u)]^{1-2\theta} P \} \right\}. \quad (16)$$

Notice that here, in contrast with Eq. (15), θ only appears in the diffusion term, and not in the drift one. The demonstration of the equivalence between Eqs. (15) and (16) is straightforward and can be found in Appendix, Sec. 2.

Let us now consider a family of models represented by Langevin equations of the type

$$\dot{u} = f(u) + g(u) \xi(t) + \eta(t), \quad (17)$$

where $\xi(t)$ and $\eta(t)$ are uncorrelated zero-mean Gaussian white noises with autocorrelation function given by

$$\langle \xi(t) \xi(t') \rangle = 2M \delta(t-t'), \quad \langle \eta(t) \eta(t') \rangle = 2A \delta(t-t'), \quad (18)$$

where $M \geq 0$ and $A > 0$. Equation (17) can be rewritten as

$$\dot{u} = f(u) + \tilde{g}(u) \zeta(t), \quad (19)$$

where the additive and multiplicative noise terms were replaced with an effective multiplicative noise given by $\tilde{g}(u) = \sqrt{\{M[g(u)]^2 + A\}/C}$ and $\zeta(t)$ is a zero-mean Gaussian white noise with autocorrelation function given by $\langle \zeta(t) \zeta(t') \rangle = 2C \delta(t-t')$ ($C > 0$). A possible demonstration of

this relation is given in the Appendix. Notice that Eq. (19) is of the form of Eq. (6) with $dW(t)=\zeta(t)dt$. As we want to analyze the influences of the additive and multiplicative noises separately, we will use Eq. (17) instead of Eq. (19) henceforth.

Following the same procedure adopted to calculate the coefficients of the FPE associated with Eq. (6), we can easily calculate the Kramers-Moyal coefficients for Eq. (17). They are

$$K_{\theta}(u) = f(u) + 2\theta M g(u)g'(u), \tag{20}$$

$$D(u) = A + M[g(u)]^2. \tag{21}$$

These results lead us to the following FPE:

$$\begin{aligned} \frac{\partial P(u,t)}{\partial t} = & -\frac{\partial}{\partial u}\{[f(u) + 2\theta M g(u)g'(u)]P(u,t)\} \\ & + M\frac{\partial^2}{\partial u^2}\{[g(u)]^2 P(u,t)\} + A\frac{\partial^2}{\partial u^2}P(u,t), \end{aligned} \tag{22}$$

which is basically Eq. (15) with an additional term due to additive noise.

For $f(u)$ derived from a potential-like function $V(u) = (\tau/2)[g(u)]^2$, the stationary solution has the form $P(u, \infty) \propto e_q^{-\beta V(u)}$ [41], where the q -exponential function is defined as follows: $e_q^x \equiv [1 + (1-q)x]_+^{1/(1-q)}$, with $[z]_+ = z$ if $z > 0$, and zero otherwise; $e_1^x = e^x$. If we consider the simple case $g(u) \propto u$, then the stationary-state distribution is a q Gaussian, i.e., $P(u, \infty) \propto e_q^{-\beta u^2}$, where q and β are given by

$$\beta = \frac{\tau + 2M(1 - \theta)}{2A}, \tag{23}$$

$$q = \frac{\tau + 2M(2 - \theta)}{\tau + 2M(1 - \theta)}. \tag{24}$$

We can verify that these results unify those obtained in [41], more specifically $q = \frac{\tau+4M}{\tau+2M}$ for the Itô case ($\theta=0$) and $q = \frac{\tau+3M}{\tau+M}$ for the Stratonovich case ($\theta=1/2$).

Let us remind at this point that the q -Gaussian form precisely is the one which, under appropriate constraints, extremizes the entropy

$$S_q[p] \equiv k_B \frac{1 - \int dx [p(x)]^q}{q - 1}, \tag{25}$$

where $S_1 \equiv S_{BG} = -k_B \int dx p(x) \ln p(x)$. For further connections between the structure of Fokker-Planck-like equations and entropy, as well as the validity of the H theorem, see [42,43].

III. GENERALIZED MOMENTS AND KURTOSIS

As already mentioned, the q Gaussian is the distribution form which extremizes entropy (25) under appropriate constraints. There are many theoretical reasons suggesting that,

in the extremization of S_q , it is convenient to express the constraints that are being imposed in an escort mean value form. Escort mean values (or q moments) are useful tools to analyze power-law distributions that frequently appear in the study of complex systems. They are defined as

$$\langle A(x) \rangle_q = \int_{-\infty}^{+\infty} A(x) f_q(x) dx, \tag{26}$$

where the *escort probability density* is given by

$$f_q(x) = \frac{[f(x)]^q}{\int_{-\infty}^{+\infty} [f(x)]^q dx}. \tag{27}$$

The characterization of a probability density in terms of its escort mean values is a natural extension of the well-known characterization of a distribution in terms of its standard moments if all of them are *finite*. The physical interpretation of q -generalized mean values demands an explanation. The quantities that are physically important in order to characterize a distribution are, for instance, the most probable value, the range of the typical values, the width of the distribution, its asymmetry, and so on. Such information is conveniently contained in the set of successive mean values of the distribution *as long as they are finite*. What can be done whenever all moments above a given one diverge? For example, if we are dealing with the Cauchy-Lorentz distribution, how can we characterize its width (obviously finite for any given such distribution)? Certainly not through its second moment since it diverges. By appropriately choosing the value of q (see below), its width can be characterized by its q variance (i.e., its variance with the escort distribution), which will also be finite.

In [44], the correct set of all escort mean values is shown, together with the set of all associated normalizing quantities that characterize a given probability density $f(x)$, even if it decays as slowly as a power law. Based on a generalization of the Fourier transform, namely, the q -Fourier transform, defined by [45]

$$\mathcal{F}_q[f](\xi) = \int_{-\infty}^{+\infty} f(x) e_q^{i\xi[f(x)]^{q-1}} dx \quad (q \geq 1), \tag{28}$$

it is possible to expand the q -characteristic function and obtain the correct exponents values necessary to perform the calculations of the q moment. The family of escort mean values which arises from this procedure is given by

$$\langle x^n \rangle_{q_n} = \frac{\int_{-\infty}^{+\infty} x^n [f(x)]^{q_n} dx}{\int_{-\infty}^{+\infty} [f(x)]^{q_n} dx}, \tag{29}$$

where

$$q_n = 1 + n(q - 1). \tag{30}$$

The kurtosis, usually defined in terms of the ratio between the fourth ($n=4$) and three times the squared second moment

($n=2$) of a given distribution, is regarded as a measure of how different a given distribution is from a Gaussian. Therefore the q kurtosis, defined in terms of q moments, is the analogous measure for q Gaussians. Its mathematical forms are defined, respectively, as follows:

$$\kappa_q = \frac{\int_{-\infty}^{+\infty} x^4 [f(x)]^{4q-3} dx}{3 \left\{ \int_{-\infty}^{+\infty} x^2 [f(x)]^{2q-1} dx \right\}^2}. \quad (31)$$

Notice that the standard kurtosis can be regarded as a particular case ($q=1$) of Eq. (31). Notice also that the number 3 in the denominator of Eq. (31) plays no special role and can equally well be replaced with the q -dependent value exactly corresponding to q Gaussians, instead of that for Gaussians.

To calculate the correct value of q we shall now consider a typical situation arising in complex systems, such as those commonly addressed within q thermostatics. Usually these systems asymptotically behave as power-law probability densities, i.e.,

$$f(x) \sim |x|^{-\gamma} \quad (|x| \rightarrow \infty, \gamma > 1). \quad (32)$$

It is clear that, if $f(x)$ is not defined on a bounded interval, the standard linear moments $\langle x^n \rangle$ may diverge above some value of n . Therefore, the standard procedure to characterize the probability density through all its moments cannot be implemented. However, in [44] it is shown how the escort mean values can overcome such difficulties. In this work, the relation is established between the exponent γ of the power-law distribution and the escort mean value parameter q . It is given by

$$q = 1 + \frac{1}{\gamma}. \quad (33)$$

So, once we know the power-law exponent γ , we will have the correct value of q , and hence the values q_n which enable the calculation of the escort mean values of the probability density.

For the special case where $f(x) = e_Q^{-\beta x^2}$ (a Q Gaussian), we have

$$e_Q^{-\beta x^2} \sim |x|^{-2/(Q-1)} \quad (|x| \rightarrow \infty). \quad (34)$$

Hence, $\gamma = 2/(Q-1)$, for $Q > 1$. In this case, relation (33) becomes

$$q - 1 = \frac{Q - 1}{2}. \quad (35)$$

A Q Gaussian is normalizable for $Q < 3$; for $Q < 1$ it has a compact support and an unbounded support for $1 \leq Q < 3$. Its second and fourth moments diverge for $5/3 < Q < 3$ and $7/5 < Q < 3$, respectively. But their second and fourth Q moments are finite for $Q < 3$. These two Q moments can

be analytically calculated [44,46] and written in terms of gamma functions,

$$\langle x^2 \rangle_Q = \begin{cases} \frac{\beta^{-1} \Gamma\left(\frac{Q}{Q-1} - \frac{3}{2}\right)}{2(Q-1) \Gamma\left(\frac{Q}{Q-1} - \frac{1}{2}\right)}, & \text{for } 1 < Q < 3 \\ \frac{\beta^{-1}}{2}, & \text{for } Q = 1 \\ \frac{\beta^{-1} \Gamma\left(\frac{Q}{1-Q} + \frac{3}{2}\right)}{2(1-Q) \Gamma\left(\frac{Q}{1-Q} + \frac{5}{2}\right)}, & \text{for } 0 < Q < 1, \end{cases} \quad (36)$$

$$\langle x^4 \rangle_{2Q-1} = \begin{cases} \frac{3\beta^{-2} \Gamma\left(\frac{Q}{Q-1} - \frac{3}{2}\right)}{4(Q-1)^2 \Gamma\left(\frac{Q}{Q-1} + \frac{1}{2}\right)}, & \text{for } 1 < Q < 3 \\ \frac{3\beta^{-1}}{4}, & \text{for } Q = 1 \\ \frac{3\beta^{-2} \Gamma\left(\frac{Q}{1-Q} + \frac{1}{2}\right)}{4(1-Q)^2 \Gamma\left(\frac{Q}{1-Q} + \frac{5}{2}\right)}, & \text{for } 0 < Q < 1. \end{cases} \quad (37)$$

We intend to compare the temporal evolution of the distributions given by the FPE of the previous section to its stationary states, which we already know to be Q Gaussians. To do it, it is convenient to normalize the Q kurtosis by its stationary value, $\kappa_Q(\infty) = \lim_{t \rightarrow \infty} \kappa_Q(t)$, which is given by Eqs. (36) and (37),

$$\kappa_Q(\infty) = \begin{cases} \frac{\Gamma^2\left(\frac{Q}{Q-1} - \frac{1}{2}\right)}{\Gamma\left(\frac{Q}{Q-1} + \frac{1}{2}\right) \Gamma\left(\frac{Q}{Q-1} - \frac{3}{2}\right)}, & \text{for } 1 < Q < 3 \\ 1, & \text{for } Q = 1 \\ \frac{\Gamma^2\left(\frac{Q}{1-Q} + \frac{1}{2}\right) \Gamma\left(\frac{Q}{1-Q} + \frac{5}{2}\right)}{\Gamma^2\left(\frac{Q}{1-Q} + \frac{3}{2}\right)}, & \text{for } 0 < Q < 1. \end{cases} \quad (38)$$

In Fig. 1 we can see the temporal evolution of the normalized Q kurtosis obtained by the numerical integration of the FPE. In the left column, we show the temporal evolution for

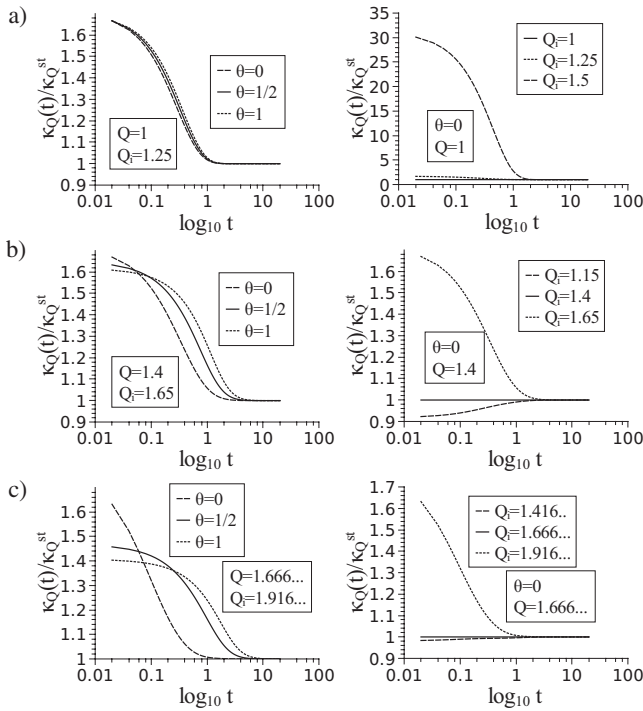


FIG. 1. On the left column are the temporal evolution of Q kurtosis for Itô ($\theta=0$), Stratonovich ($\theta=1/2$), and backward Itô ($\theta=1$) stochastic procedures. The numerical results are normalized by its stationary value $\kappa_Q^{st}=\kappa_Q(\infty)$. In (a) we have $Q=1$ and $Q_i=5/4$, (b) $Q=7/5$ and $Q_i=1.65$, and (c) $Q=1.666\dots$ and $Q_i=1.9166\dots$. As expected, for sufficient large times the numerical results approach unity. The right column shows the temporal evolution of the normalized Q kurtosis with a fixed θ . In (a) we have $Q=1$ and $Q_i=1, 5/4$, and $3/2$; (b) shows the results for $Q=7/5$ and $Q_i=1.15, 7/5$, and 1.65 ; and (c) the results for $Q=5/3$ and $Q_i=1.4166\dots, 5/3$, and $1.9166\dots$.

the three stochastic procedures above mentioned: Itô ($\theta=0$), Stratonovich ($\theta=1/2$), and backward Itô ($\theta=1$). In the right column we show the same temporal evolution of the normalized Q kurtosis but now with $\theta=0$ and different values of Q_i . As expected, the kurtosis approaches monotonically its stationary value for all values of parameter Q_i .

In Fig. 2 we can compare the behavior of the standard kurtosis and the Q kurtosis given by Eq. (38). As we can see, the Q kurtosis diverges at $Q=-1$, but is finite for $-1 < Q < 3$, while the standard kurtosis diverges for $7/5 < Q < 3$. As expected, the Q kurtosis coincides with the standard one for $Q=1$. Furthermore, the Q kurtosis is a monotonically decreasing function of Q . Based on this feature, experimentalists could use it as a calibration curve that allows the determination of the proper value of the entropic index Q for a given system.

IV. CONCLUSIONS

Based on stochastic integrals which unify the Itô, Stratonovich, and kinetic approaches, we obtain two different [namely, Eqs. (15) and (16)], although equivalent, forms of a unified Fokker-Planck equation. Form (15) recovers, as par-

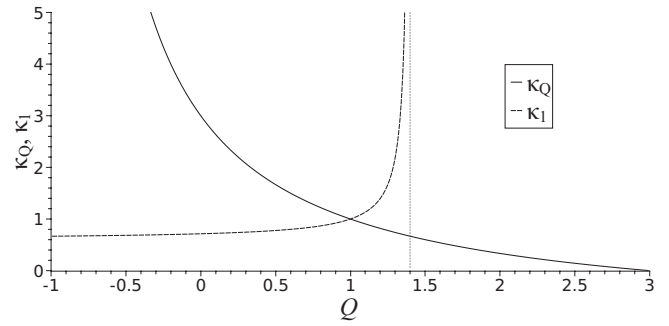


FIG. 2. The standard kurtosis κ_1 and the Q kurtosis κ_Q as functions of Q . We verify that both have the same value of 1 for $Q=1$. For $Q > 1$, the use of κ_Q is generically more convenient than using κ_1 . Indeed, the latter one diverges at $Q=7/5$, while κ_Q remains finite up to the maximal value $Q=3$. In contrast, for $Q < 1$, κ_1 is more convenient than κ_Q . Indeed, κ_Q diverges at $Q=-1$, whereas κ_1 remains finite down to $Q \rightarrow -\infty$. In the $Q \rightarrow -\infty$ limit, κ_1 saturates at the value of $3/5$.

ticular instances, equations currently available in the literature. Form (16) is herein introduced. Both Fokker-Planck forms exhibit an explicit dependence on the unifying parameter θ . One of these forms [Eq. (15)] exhibits this dependence only in the drift term, which [see Eq. (13)] is in turn composed by a deterministic part and a noise-induced part. In the other form [Eq. (16)], θ only appears in the diffusion term. It appears therefore that, in general, the noise-induced drift is equivalent to a nonlinear diffusion term (only for the particular values $\theta=0$ and $\theta=1/2$ the diffusion coefficient is linear).

We present, based on escort mean values, an explicit form for the Q -generalized kurtosis to study the convergence towards the Q -Gaussian stationary solution. Paper [46] focuses on the porous-medium equation (homogeneous nonlinear Fokker-Planck equation). In the present paper we deal with another interesting equation, namely, a linear inhomogeneous one, whose stationary state also has the Q -Gaussian form, which once again allows us to calculate explicit forms for the Q kurtosis. As expected, the H theorem also holds for the linear inhomogeneous Fokker-Planck equation, and this can be verified by Eq. (24) in Ref. [43].

In order to follow the time evolution toward the Q -Gaussian stationary states, we have evaluated the approach of the Q kurtosis to its stationary-state limits for different values of θ . The difference in this convergence (see Fig. 1) for the same q Gaussian with different θ 's is due to a change in the multiplicative noise amplitude needed to compensate the change in the θ value [see Eq. (24)].

Finally, we propose that, if a system is well described by nonextensive statistical mechanics and has a Q -Gaussian form for its probability density function, the Q kurtosis can be used as a calibration curve to determine, from data, the best value of the entropic index Q .

ACKNOWLEDGMENTS

We have benefited from useful remarks from L. Borland. We acknowledge partial financial support from the Brazilian

agencies Conselho Nacional de Desenvolvimento Científico e Tecnológico (CNPq) and Fundação Carlos Chagas Filho de Amparo à Pesquisa do Estado do Rio de Janeiro (Faperj).

APPENDIX

1. Integral $I_\theta[W(t)] = \int_{t_0}^t dW(t')W(t')$

Here, we perform, in detail, the calculation of the stochastic integral of the function $G(t) = W(t)$ in terms of discrete sums. The definition of the unified stochastic integral is

$$I_\theta[G(t)] \equiv \text{ms-lim}_{n \rightarrow \infty} S_n = \text{ms-lim}_{n \rightarrow \infty} \sum_{i=1}^n [\theta G_i + (1 - \theta)G_{i-1}] \Delta W_i. \quad (\text{A1})$$

If $G(t) = W(t)$, we have

$$S_n = \sum_{i=1}^n [\theta W_i \Delta W_i + (1 - \theta)W_{i-1} \Delta W_i]. \quad (\text{A2})$$

The terms $W_i \Delta W_i$ and $W_{i-1} \Delta W_i$ can be written as

$$2W_i \Delta W_i = (W_i)^2 + (\Delta W_i)^2 - \overbrace{(W_i - \Delta W_i)^2}^{W_{i-1}}, \quad (\text{A3})$$

$$2W_{i-1} \Delta W_i = \underbrace{(W_{i-1} + \Delta W_i)^2}_{W_i} - (W_{i-1})^2 - (\Delta W_i)^2. \quad (\text{A4})$$

Notice that in both equations we have terms like $(W_i^2 - W_{i-1}^2)$ under a summation. Therefore, only the first and the last terms need to be taken into account in the summation.

Indeed, all the others mutually cancel, and we finally have

$$\sum_{i=1}^n W_i \Delta W_i = \frac{W^2(t) - W^2(t_0) + \sum_{i=1}^n (\Delta W_i)^2}{2}, \quad (\text{A5})$$

$$\sum_{i=1}^n W_{i-1} \Delta W_i = \frac{W^2(t) - W^2(t_0) - \sum_{i=1}^n (\Delta W_i)^2}{2}. \quad (\text{A6})$$

Taking this last two equations into account and the fact that $\text{ms-lim}_{n \rightarrow \infty} \sum_{i=1}^n (\Delta W_i)^2 = t - t_0$, we obtain

$$I_\theta[W(t)] \equiv \text{ms-lim}_{n \rightarrow \infty} S_n = \frac{1}{2} \{ [W(t)]^2 - [W(t_0)]^2 - (1 - 2\theta)(t - t_0) \}, \quad (\text{A7})$$

which is the same result presented in Eq. (5).

2. How to get Eq. (15) from Eq. (16)

In Sec. II we claim that Eqs. (15) and (16) are equivalent. To prove this statement, we only need to show the following equality:

$$K_0 P - \frac{1}{2} [D]^{2\theta} \frac{\partial}{\partial u} \{ [D]^{1-2\theta} P \} = K_\theta P - \frac{1}{2} \frac{\partial}{\partial u} [DP],$$

where the dependence on the dynamical variable u was omitted. To do this, we perform the derivative on the left-hand side:

$$D^{2\theta} \frac{\partial}{\partial u} [D^{1-2\theta} P] = D^{2\theta} \left[(1 - 2\theta) D^{-2\theta} P \frac{\partial D}{\partial u} + D^{1-2\theta} \frac{\partial P}{\partial u} \right] = -2\theta \frac{\partial D}{\partial u} P + \left[P \frac{\partial D}{\partial u} + D \frac{\partial P}{\partial u} \right] = -2\theta \frac{\partial D}{\partial u} P + \frac{\partial}{\partial u} [DP]. \quad (\text{A8})$$

So,

$$K_0 P - \frac{1}{2} [D]^{2\theta} \frac{\partial}{\partial u} \{ [D]^{1-2\theta} P \} = K_0 P - \frac{1}{2} \left\{ -2\theta \frac{\partial D}{\partial u} P + \frac{\partial}{\partial u} [DP] \right\} = K_0 P + \theta \frac{\partial D}{\partial u} P - \frac{1}{2} \frac{\partial}{\partial u} [DP], \quad (\text{A9})$$

where $K_0 P + \theta \partial_u DP = K_\theta$.

3. Equivalence between $\dot{u} = f(u) + g(u)\xi(t) + \eta(t)$ and $\dot{u} = f(u) + \tilde{g}(u)\xi(t)$

Suppose two different Langevin equations for the stochastic processes $u(t)$, both with the same deterministic drift force $f(u)$ but with different noise sources,

$$\dot{u} = f(u) + g(u)\xi(t) + \eta(t), \quad (\text{A10})$$

$$\dot{u} = f(u) + \tilde{g}(u)\xi(t). \quad (\text{A11})$$

The mean-value solution of both equations is

$$\langle u(t) \rangle = \int_0^t dt' f(u). \quad (\text{A12})$$

Let us calculate the variance corresponding to both equations. From the first one we obtain

$$\langle [u(t) - \langle u(t) \rangle]^2 \rangle = \left\langle \int_0^t dt' \int_0^t dt'' \{g[u(t')] \xi(t') + \eta(t')\} \{g[u(t'')] \xi(t'') + \eta(t'')\} \right\rangle \quad (\text{A13})$$

$$= \int_0^t dt' \int_0^t dt'' \{g[u(t')] g[u(t'')] \underbrace{\langle \xi(t') \xi(t'') \rangle}_{=2M \delta(t'-t'')} + \underbrace{\langle \eta(t') \eta(t'') \rangle}_{=2A \delta(t'-t'')} \} \quad (\text{A14})$$

$$+ g[u(t')] \underbrace{\langle \xi(t') \eta(t'') \rangle}_{=0} + g[u(t'')] \underbrace{\langle \xi(t'') \eta(t') \rangle}_{=0} \quad (\text{A15})$$

$$= \int_0^t dt' (2M \{g[u(t')]\}^2 + 2A). \quad (\text{A16})$$

From the second one we obtain

$$\langle [u(t) - \langle u(t) \rangle]^2 \rangle = \int_0^t dt' \int_0^t dt'' \tilde{g}[u(t')] \tilde{g}[u(t'')] \underbrace{\langle \zeta(t') \zeta(t'') \rangle}_{2C \delta(t'-t'')} \quad (\text{A17})$$

$$= \int_0^t dt' 2C \{\tilde{g}[u(t')]\}^2. \quad (\text{A18})$$

If we want that both Langevin equations give us the same mean-squared value for the stochastic process $u(t)$, we must match the integrands of Eqs. (A13) and (A17). This leads to the following relation among noise terms:

$$\tilde{g}(u) = \sqrt{\frac{M[g(u)]^2 + A}{C}}. \quad (\text{A19})$$

This result shows that equations like Eq. (17) can be rewritten as in Eq. (19), with an effective multiplicative noise term given by Eq. (A19).

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